
Technical Report*

Neuronal Rhythmicity and its Rate Variability[†]

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Abstract

Rhythmic surface Electromyographic (sEMG) bursts for which their returns occur at irregular times are analyzed and compared among healthy and injured central nervous systems. The rhythmic electrophysiological activity in the present study has been defined as “*Bursting Rate Variability*” (BRV) due to the inherent aperiodicity of successive sEMG bursts, which are concomitant with “doublet” waveforms in the D8 subband of the Daubechies 3 wavelet decomposition of the raw signal. A key element in the analysis is the precise time-localization of D8 doublets that requires a statistical waveform matching between the D8 doublet and the burst in the raw sEMG signal. This study has been conducted over a period of 10 years, in which 7 healthy and 2 unhealthy individuals volunteered and presented a total of ~8,000 doublets: it was observed that doublets were more prevalent in healthy than unhealthy subjects, and that the probability distribution of return times was best fitted with Normal Mixtures in healthy subjects, compared to the Weibull distribution as the best fit in unhealthy ones based on the corrected Akaike Information Criterion (AICc) for model selection. Finally, the rate in the occurrence of doublets appears to be within 60-88 bursts per minute, suggesting a possible connection between BRV and the heart rate dynamics.

Keywords: Bursting, Central Pattern Generator, Discrete Wavelet Transform, Coherence at a distance, Synchronization, Multi-channel surface electromyography, Return time, Weibull distribution, Normal Mixture distributions

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1 Introduction

The surface electromyographic (sEMG) activity recorded along the paraspinal muscles has been seen to present standing wave properties [1, 2] when the research subject is placed in the prone position and is applied light pressure at some specific “gateway” points of the spine (usually the neck and the coccyx) to elicit the oscillation.

The motion usually starts in a chaotic fashion at the distal ends of the spine, propagates caudally, until it settles in a standing wave pattern, which can undergo “period halving bifurcations,” transitioning away from chaos [3]. At that stage, no further digital stimulus is required, indicating that the rhythmic movement is innervated by a Central Pattern Generator (CPG) [4, 5], as argued in [1]. A further confirmation of the CPG hypothesis is that two quadriplegic subjects have been able to sustain the so-called *spinal wave* [1, 6, 7].

Besides CPG, another important aspect of this movement is *coherence at a distance*; the antinodes of the wave are indeed in coherent motion, with a wavelength in the order of ~ 1 m, hence qualifying as *coherence at a distance* [8, 9, 10].

The specificity of this CPG-innervated, coherent motion is confirmed by the Daubechies 3 (db3) wavelet decomposition [11, 12] of the sEMG signal, more specifically, by the “doublet” waveforms repeating themselves in an aperiodic fashion in the D8 subband of the wavelet decomposition. Due to surface electromyography being a superposition of multiple Motor Unit Action Potential (MUAP) trains (see [13, Fig. 3-4]), it is still a bit unclear at this stage whether the observed “doublets” are two discharges of motor units firing in unison, or single discharges of coupled motor units firing one after the other (see [14, 15, 16] for further details on the double discharge phenomenon). In either case, the observed doublet waveform conforms to the definition of an “*exceptional doublet*” of [14, Fig. 5B], “exceptional” in the sense of large intradoublet interspike interval. Indeed, the intradoublet interspike interval duration of ~ 62.5 ms—that we have measured peak-to-peak from the D8 doublet waveform—exceeds the conventional limits of 2-20 ms according to the standardization of doublets set by the American Association of Neuromuscular & Electrodiagnostic Medicine [17]. However, it has been reported in [14] that the standard range of doublets can be exceeded as Piotrkiewicz et al. [14] report exceptional doublets with 37 ms of intradoublet time in the human soleus muscle. Furthermore, the Discrete Wavelet Transform (DWT) has also served to obtain the doublet total time duration [18] of ~ 125 ms measured from onset to offset of the wavelet waveform at scale 8, which spans the time comprised by two wavelet transform coefficients at this scale.

Besides these early findings, the crucial observation that launched this research is the near-synchrony between the onsets of the doublets observed on the D8 traces and the onsets of the

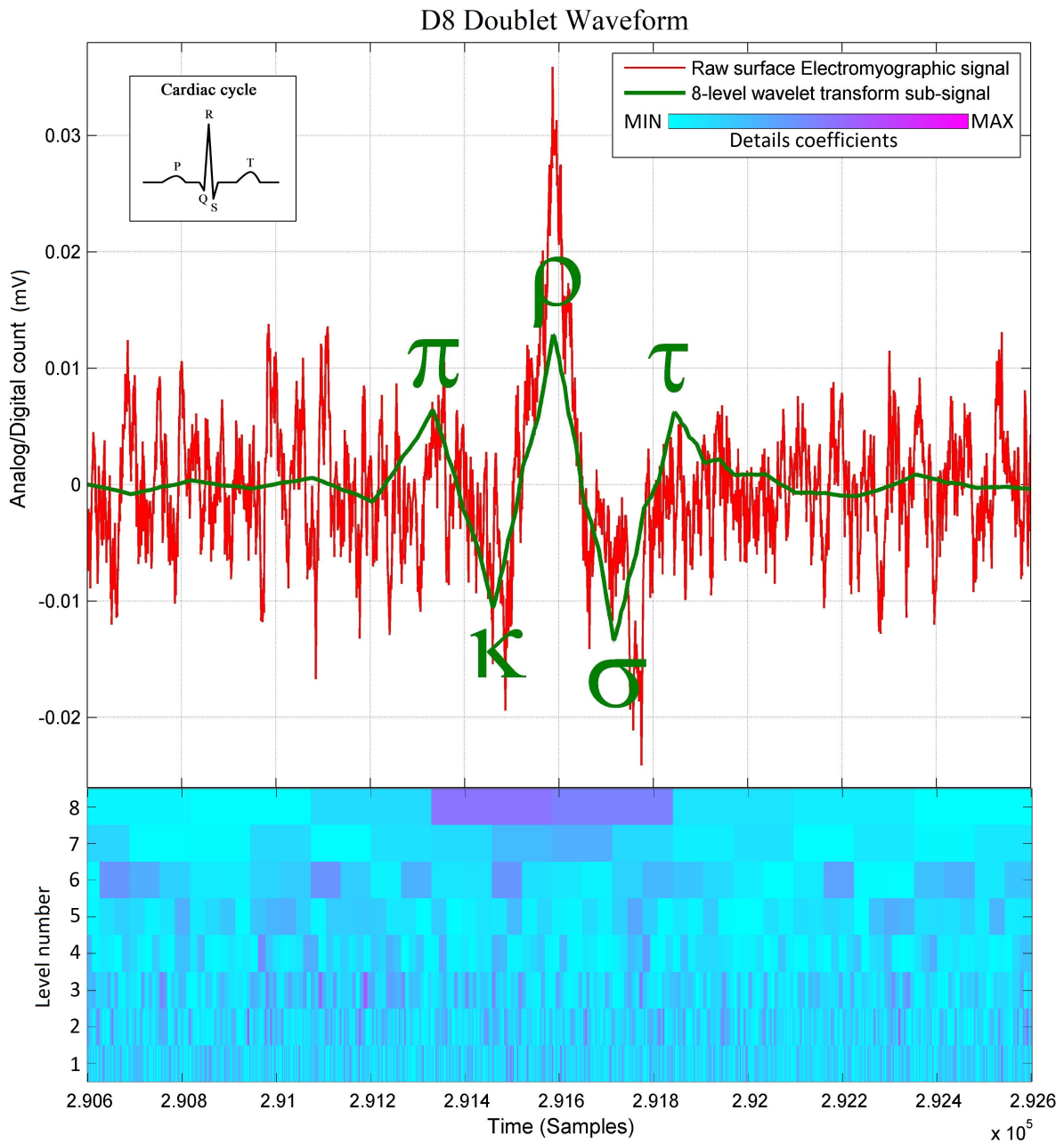


FIGURE 1

Raw paraspinal surface electromyographic signal of a healthy individual, superimposed with its 8-level wavelet transformed sub-signal; its scalogram is shown below: The D8 doublet is a precise sequence of (+) peaks and (-) dips, defined here as the π - κ - ρ - σ - τ sequence because of its similarity with the cardiac cycle. The concordance between the D8 and the burst here appears naturally without preprocessing.

bursts of accrued sEMG activity visible on the raw signal traces, as shown in Figs. 1 and 2. The “nearness” of the time localizations of the doublets and the bursts is crucial here.

By definition of the wavelet decomposition, the repetition of the DWT frame generating the D8 subband is periodic, while the sequence of bursts is not. Therefore, some time-shifting of the

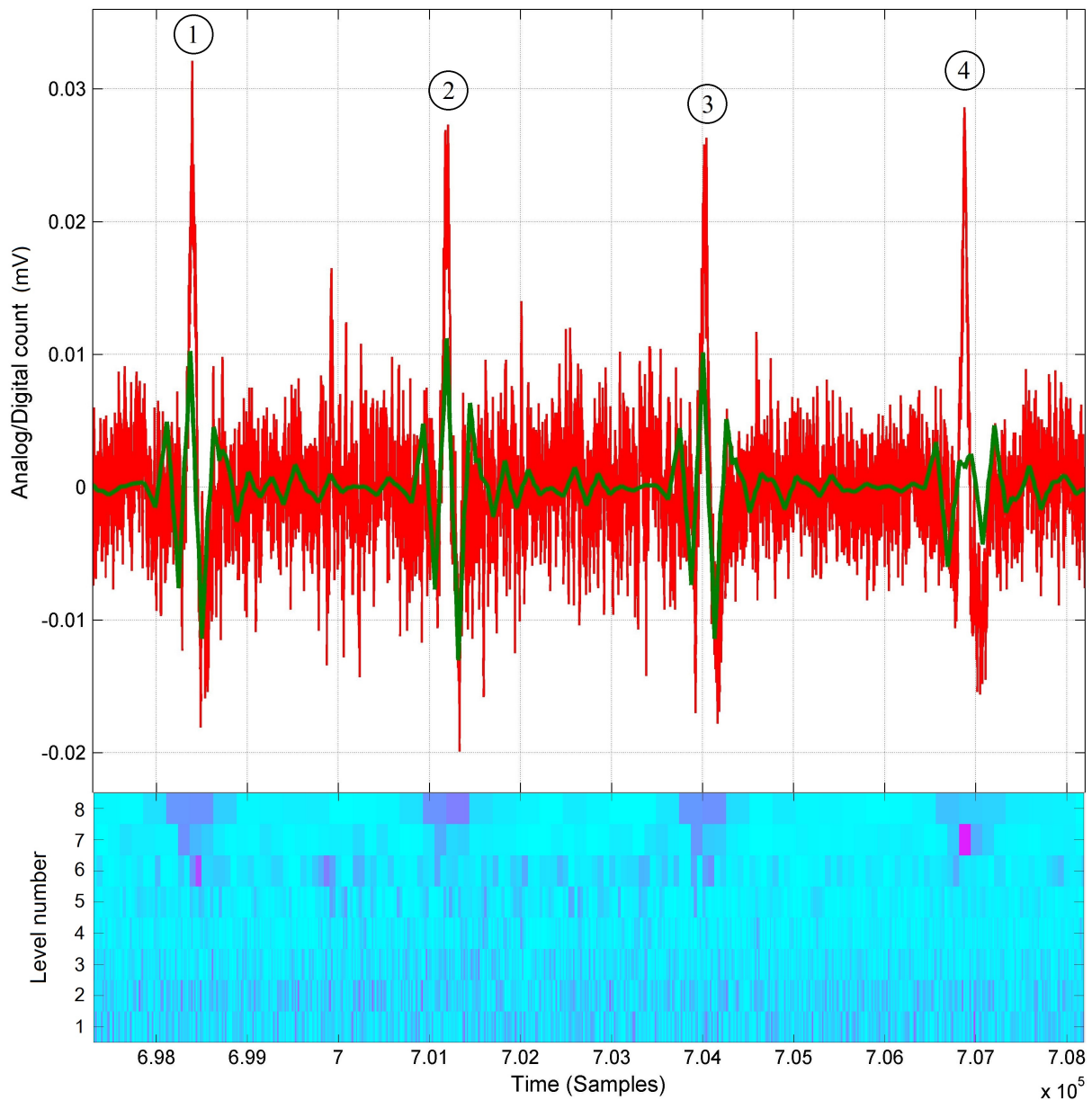


FIGURE 2

Raw thoracic surface electromyographic signal (in red) of control subject #4, superimposed with its 8-level wavelet transformed sub-signal (in green); its scalogram is shown below: Note that at the first, second, and third bursts the matching between the raw sEMG signal and its 8-level sub-signal is better than the one at the fourth burst. The recovery of the matching between the fourth burst and its D8 doublet is shown in Fig. 4, where different time delays are applied before wavelet processing.

raw signal trace is necessary to acquire a good *waveform matching* between a *specific* burst and its D8 doublet. It turns out that this time-shifting is different from one burst to the other, leading to a variability of the time interval between successive bursts, which is referred to as *Bursting Rate Variability (BRV)*.

The empirical distribution of the time intervals between successive bursts differ from one subject to another, but the types of continuous probability distributions already differentiates the BRV of quadriplegic (Figs. 8a and 8b) versus control subjects (Figs. 7a, 7b, 7c, and 7d). The quadriplegic subjects consistently presented the *Weibull* (Type III Extreme Value [19]) distribution as the best fit, and *mixtures of normal* distributions in the case of control subjects, this discrepancy already points to some neurophysiological applications of BRV, but probably more importantly, it could allow for a *dynamical* understanding of the bursting or doublet phenomenon in *abnormal* conditions. The various sEMG signals are indeed generated by some complicated dynamics, nearly impossible to model from “first principles.” However, the Generalized Extreme Value (GEV) theory [20] endeavors to identify the *qualitative* properties of such dynamics, no matter how complicated, from the return time of extreme events, like the bursts.

2 Methods

2.1 Methods: control and quadriplegic subjects

For our analysis, a population of 9 volunteers, 7 control (healthy) and 2 quadriplegic subjects (presenting a total of $\sim 8,000$ *doublets*) were chosen, the latter subjects with a cervical spinal cord injury at the C5 vertebral level [1, 7].

To draw an objective comparison between the two types of subjects, the subjects with quadriplegia and the healthy ones had their recordings taken during the same session. Before recordings, the subjects had signed the Informed Consent drafted by the investigators and approved by the University Park Campus (UPC) Institutional Review Board (IRB) of the University of Southern California.

2.2 Methods: electrode placement

The data utilized in this investigation have been recorded over a period of a little more than 10 years. All along those recordings, we have followed a consistent recording protocol: Surface electromyography (sEMG) reduced-noise tripolar electrodes (“Uni Patch Tyco EMG Electrodes Round Disk 7500 2.25 diameter Ag snaps”) were placed at cervical (C2-C3), thoracic (T4-T6), lumbar (L3), and sacral (S2-S4) positions; all with the same sampling rate of 4,000 samples per second.

The sensitive input prongs of the front-end electronics were initially at a 45-deg. angle with the muscle fibers and subsequently aligned with the back-muscle fibers, without significant difference observed in the results.

2.3 Methods: equipment

The most recent (<4 years) recordings were made with an Insight Discover sEMG station together with a Measurement ComputingTM USB-1608FS device for analog-to-digital conversion, while the earlier recordings (10 years ago) were done with an Insight Millennium sEMG station interfaced with a Computer Board PCMCIA DAS16/16 card analog-to-digital converter.

2.4 Methods: wavelet transform

We picked up the Daubechies DB3 wavelet decomposition, originally [1] for the motivation that its D8 subband provided the best correlations among the subband signals at different points along the spine (hence promoting the “coherence at a distance” aspect). Later [3], however, it was discovered that under some conditions the D7 subband was preferable.

Parallel to this line of thoughts, it was found by the present and other investigators that the mother function of the DB3 mimics the MUAPs detected by the electrodes [1, 21], which makes the DB3 the ideal tool for picking up relevant waveforms in an otherwise messy sEMG signal corrupted by noise and motion artifacts.

Furthermore, it was found that the D8 of the db3 allowed for accurate localization of doublets most of the times [1, 2, 3]. A wavelet similar to db3, the Daubechies 4 (db4), appears to mimic a doublet but did not provide satisfactory results as it failed to pick up the onset of some bursts as accurately as the db3 did it. Even though the db4 has more rounded crests and troughs, the db3 kernel still adheres more to the shape of observed electromyographic bursts and the exceptional doublets of [14, Fig. 5B] as the amplitude of the first trough in the db3 wavelet is less negative than its second trough as in [14, Fig. 5B]; however, this reverses in the case of the db4 kernel, in which the first trough is more negative than the second one. Another wavelet tested face-to-face against db3, the Daubechies 2 (db2), did not clearly show the bursts start and end points, compared to db3 that allowed better differentiation between two adjacent doublets.

2.5 Methods: waveform matching

To solve the “shift variance” issue of the DWT of many signals [22] causing in our particular application some of the D8 doublets not to have well-defined peaks and dips (see doublet #4 in Fig. 2), or causing the $\pi\kappa\rho\sigma\tau$ peaks and dips to be horizontally offset with respect to the raw sEMG (see the ρ -peak of doublets #1 and #3 in Fig. 2), we shift the sampling times of the

wavelet coefficients by delaying the time at which the DWT begins. This makes the coefficients span different sections of the same raw sEMG signal.

A matching increment of the π -wave, $\pi\kappa$ -slope, κ -wave, $\kappa\rho$ -slope, ρ -wave, $\rho\sigma$ -slope, σ -wave, $\sigma\tau$ -slope, and τ -wave with the raw sEMG burst around doublet #4 of Fig. 2 is obtained by omitting the first 20, 40, and 60 samples (see Figs. 3b, 3c, and 3d, respectively) before wavelet processing. This results in recovering the precise peak-dip sequence of the $\pi\kappa\rho\sigma\tau$ complex of doublet #4” in the raw burst signal, until achieving the benchmark waveform match of Fig. 1.

Due to the periodic nature of the DWT, continuing to delay the signal by omitting the first 80, 100, and 120 samples (see Figs. 3e, 3f, and 3g, respectively) before wavelet processing, results in getting away from and losing the previously recovered peak-dip sequence obtained with doublet #4”. This implies that the amount of delay is not necessary to be longer than one fourth of the time a single coefficient spans at this particular level, in our case, since an 8-level coefficient spans ~ 62.5 ms, then a 15 ms delay (omitting the first 60 samples) was enough to recover the $\pi\kappa\rho\sigma\tau$ peak-dip sequence of doublet #4 as shown in Fig. 3d.

As a consequence of the variability in the occurrence of the bursts, the process of delaying the raw sEMG signal before wavelet processing to achieve optimal waveform matching is different from one raw bursting waveform to another one. Fig. 4 shows how the best-suited time delay for doublet #4” is not the best suited for doublets #1”, #2”, #3”, as compared with doublets #1, #2, #3 of Fig. 2 when no time delay had been applied.

To make the above procedure optimal, the errors of this $\pi\kappa\rho\sigma\tau$ versus raw burst waveform matching are gathered in a multi-objective optimization function, of which the Pareto front is identified using expert rules. If a given delay for the $\pi\kappa\rho\sigma\tau$ waveform parameters is on the Pareto-optimal front as in Fig. 3d, then this provides a good match; whereas if some delays are not on the Pareto-optimal front as those in Figs. 3a, 3b, 3c, 3e, 3f and 3g, then they do not provide a good match.

The expert rules that automatically obtain the minimal error—illustrated in Fig. 3 with vertical lines—between the ρ -wave locations of each D8 doublet and the peak of its raw burst consist of a series of nested conditional statements (*if-then-else* rules) inside a *for* loop, with index value that represents the chronological position of each doublet appearing in the time series, (see flowchart of Fig. 5).

Due to the raw sEMG signal being corrupted with high-frequency noise, and in order to increase the precision in finding the times at local maxima for each D8 doublet, a Savitzky-Golay filter is applied to the raw sEMG before being processed by the expert system. In the present study, three pre-programmed delays of 5 ms, 10 ms, and 15 ms have been entered into

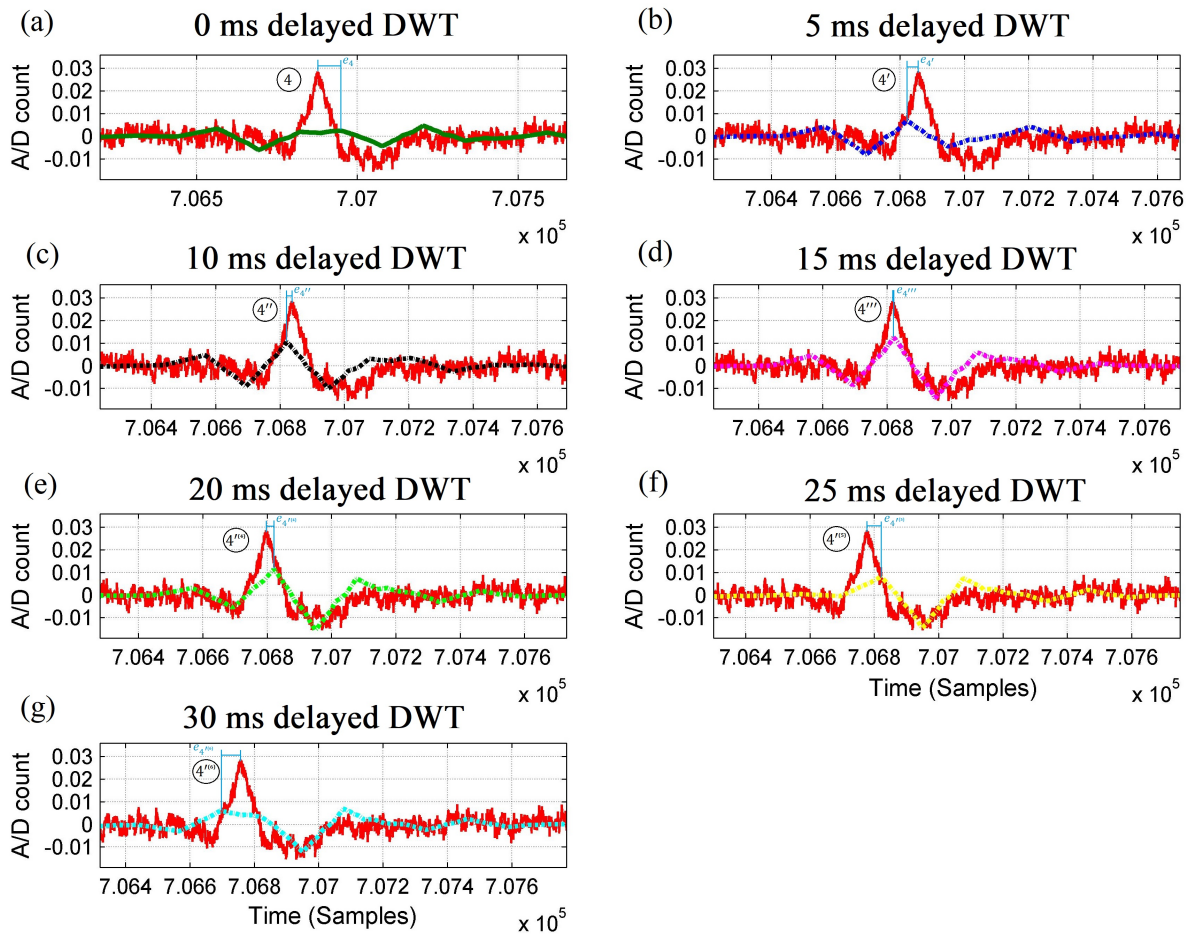


FIGURE 3

Raw thoracic surface electromyographic signal of control subject #4 (in red) centered around doublet #4 of Fig. 2, superimposed with its 8-level wavelet transformed sub-signal when (a) no delay (in dark green), (b) 5 ms delay (in dark blue), (c) 10 ms delay (in black), (d) 15 ms delay (in magenta), (e) 20 ms delay (in light green), (f) 25 ms delay (in yellow), and (g) 30 ms delay (in cyan) is applied to the raw sEMG signal before wavelet processing, where the peak time errors of the ρ -wave are $|e_{4''''}| < |e_{4'''}| < |e_{4''(4)}| < |e_{4'}| < |e_{4'(5)}| < |e_{4'(6)}| < |e_4|$: The D8 doublet with the smallest error (#4''') belongs to the Pareto-optimal front, whereas D8 doublets #4, #4', #4'', #4''(4), #4''(5), and #4''(6) do not.

the expert system algorithm as an example to attain the best waveform matching among distinct delays. Fig. 3 shows snapshots of the effect of delaying the sEMG signal every 5 ms delays. Since the DWT is periodic, a complete cycle occurs in 30 ms, thus, guidelines on pre-setting the delay values are in the range from 1 to 30 ms. It is worth noting that the expert system also considers the case of no delay time for the D8 doublets that are naturally matched with the sEMG bursts, which is represented with a zero-delay time.

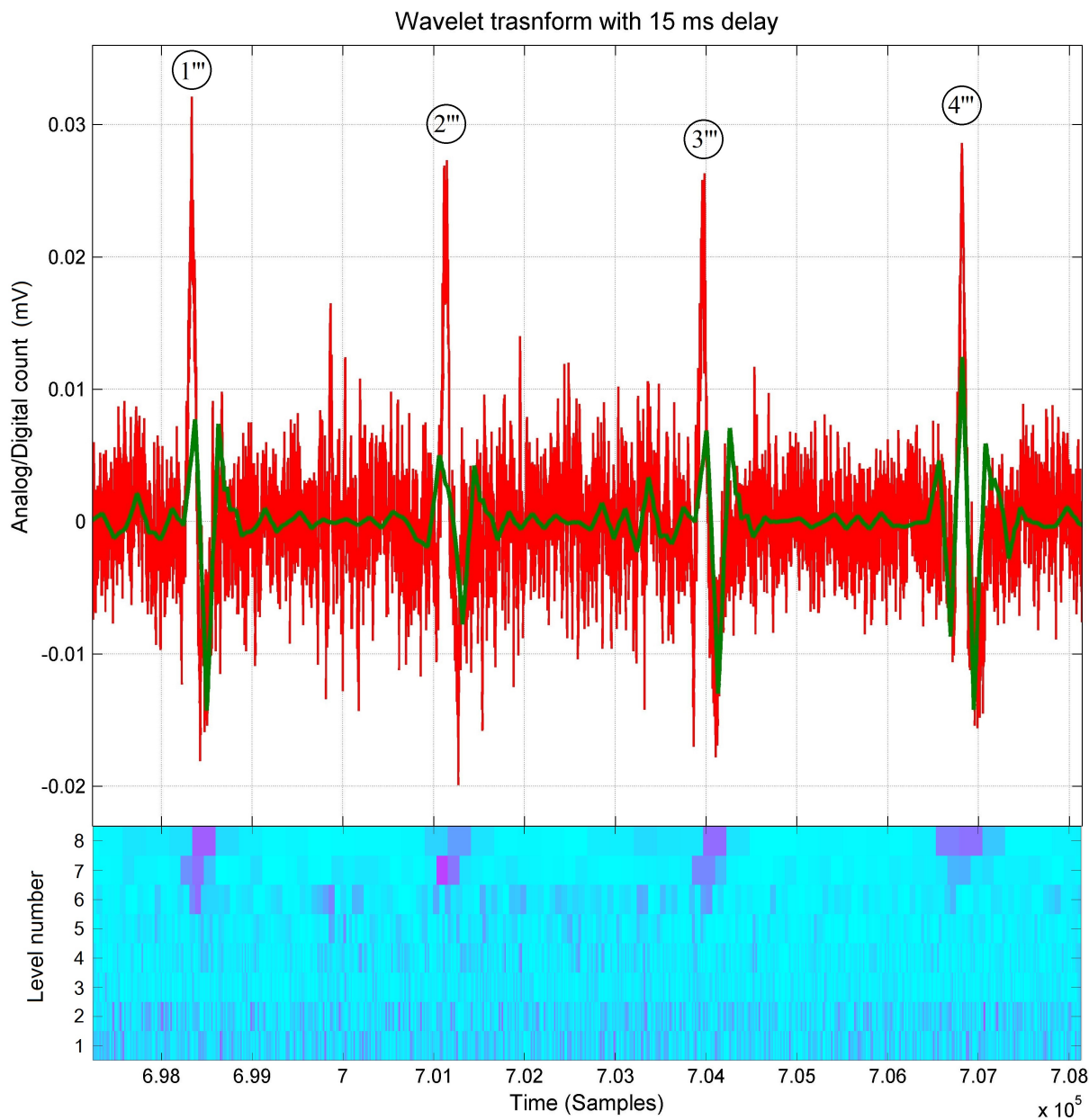


FIGURE 4

Raw thoracic surface electromyographic signal (in red) of subject #4 (control) time-shifted by 15 ms (in red), superimposed with its 8-level wavelet transformed sub-signal (in green); its scalogram is shown below: A time shift of 15 ms (skipping the first 60 samples) on the raw signal before wavelet processing is sufficient for the fourth D8 doublet to better match the raw sEMG signal than when no time shift is applied to the raw signal as in Fig. 2.

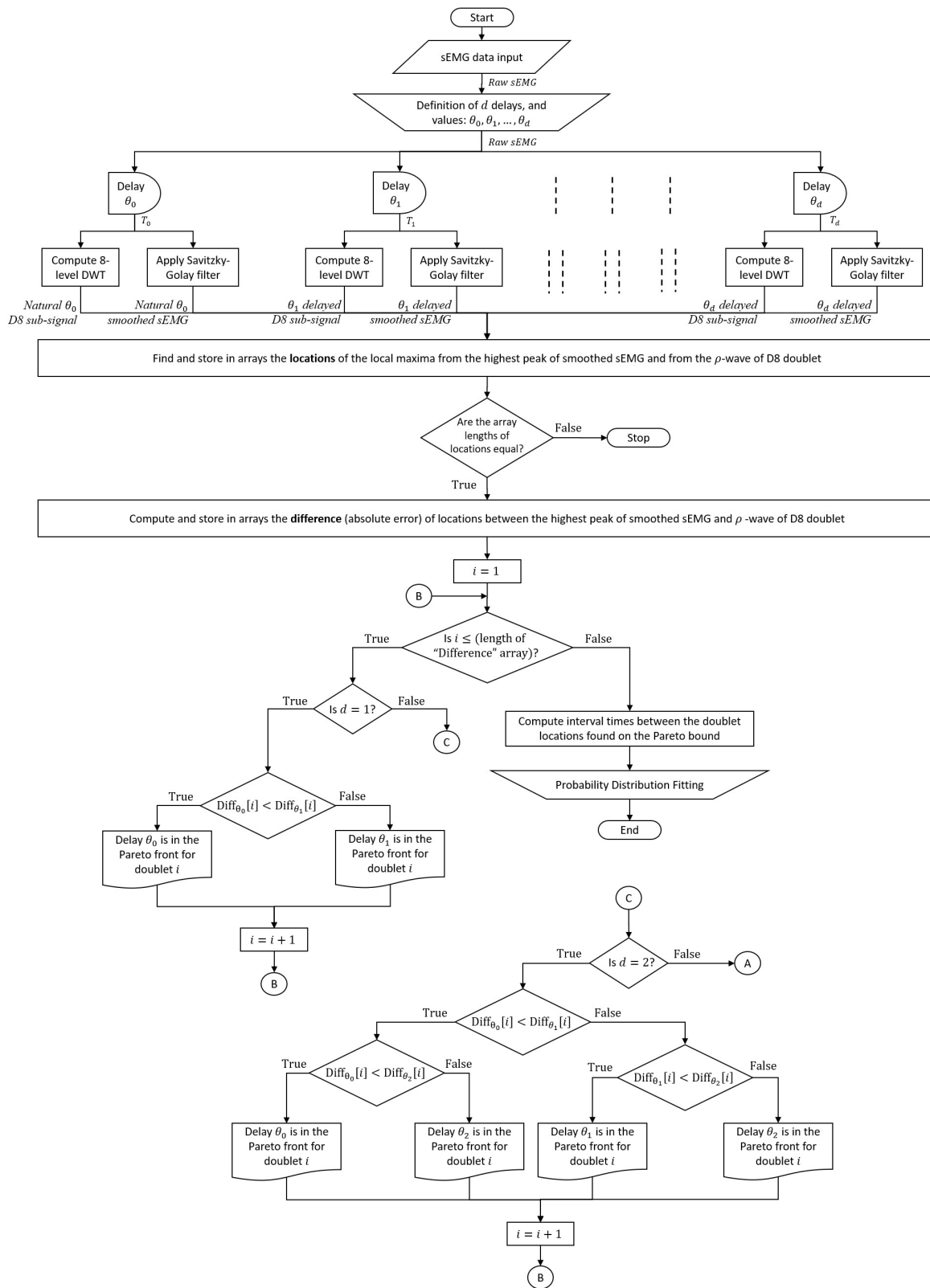
The complete sEMG signal processing technique, from data collection to probability distribution fitting, passing through 8-level DWT, delaying, smooth filtering, and including the expert rules is depicted in the flowchart of Fig. 5.

2.6 Methods: empirical vs theoretical distributions and statistical software

There are two ways to construct the histogram of the time between the ρ -waves of successive D8 waveforms, or “return time” for short: In one procedure are included in the sample set only those $\rho\rho$ -intervals between ρ -waves that are *naturally* matched with their corresponding bursts (without time shifting), discarding those that do not appear to match, such as doublet #4 of Fig. 2. However, some doublets might appear to naturally match with their sEMG bursts—for instance, doublets #1 and #3 of Fig. 2—and considering the ρ -wave locations of these two doublets would add unwanted observational errors to the distribution fit analysis since a better match was found by the expert system (see doublets #1’ and #3’ of Fig. 6). These undesired errors can be eliminated in the other “*enhanced*” procedure achieved with the expert system and the time-shifted DWT at various time delays; therefore, the latter procedure shall be preferred to increase preciseness in the distribution fitting of return times.

In the *enhanced* procedure, the histogram is constructed with accrued accuracy with $\rho\rho$ -intervals after optimal time shifting to match *all* bursts with their respective D8’s. The latter, by the same token, also increases the sample size.

The statistical software SAS[®] Studio 3.4 and JMP Pro 13 (both by the SAS Institute) were used to find the theoretical probability distributions that best fit the frequency histograms of doublet return times based on the (corrected) Akaike Information Criterion (AICc) for model selection; the results are summarized in Table 1. In the case where the Weibull distribution was the best ranked in the AICc sense, it was checked for a goodness of fit using the Cramer-von Mises-W test. As for the best-ranked normal mixture distribution, the Pearson’s chi-squared test was used. In both cases, the *null hypothesis* (H_0) states that the observed frequency distribution is consistent with the estimated theoretical distribution, and small p-values (<0.05) would reject H_0 in favor of the *alternative hypothesis* (H_1) that the data is not from the theoretical distribution.



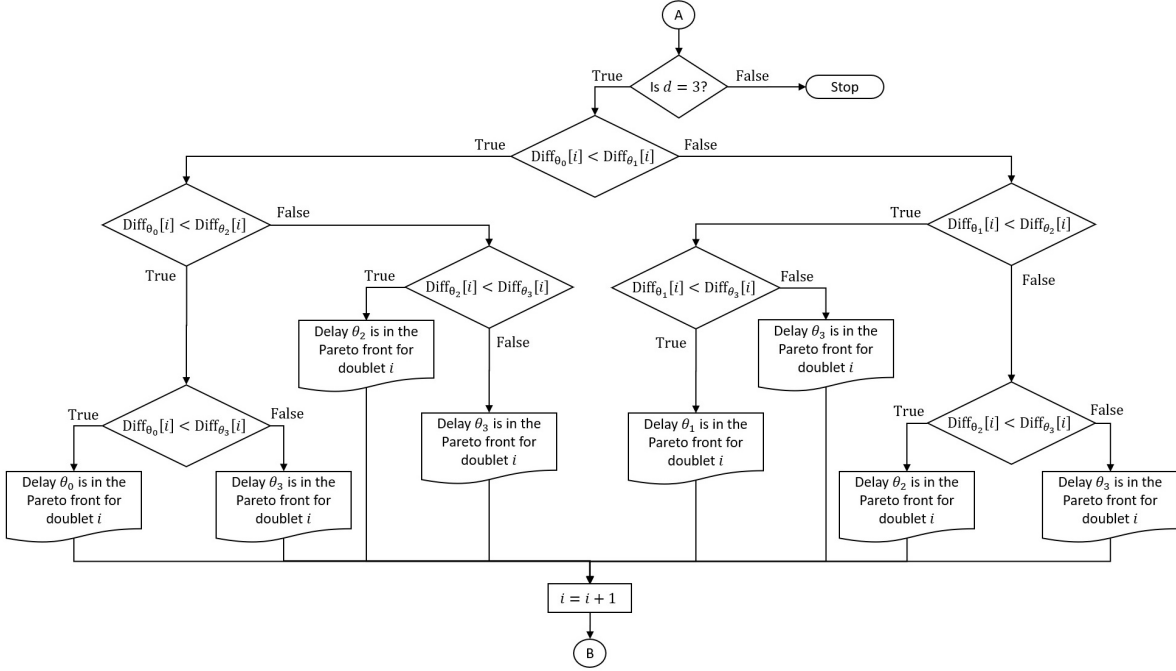


FIGURE 5

Complete sEMG signal processing and analysis flowchart: The raw sEMG signal passes through d predefined delays to time shift its 8-level sub-signal and its smoothed sEMG signal. In the example of this paper, the number of delays is $d = 3$, with values of $\theta_0 = 0$ ms, $\theta_1 = 5$ ms, $\theta_2 = 10$ ms, $\theta_3 = 15$ ms. Thus, the expert system finds the minimal errors among ρ -waves (shown in Fig. 4) between the $d + 1$ pairs of signals that come from the same sEMG trace, enhancing waveform matching and providing the Pareto front $\rho\rho$ -interval times (and/or ρ -wave magnitudes) for probability distribution analysis.

3 Model selection criterion and sample size guidelines

3.1 Best fitting probability distribution

Akaike's approach to finding the best probability distribution fit is a Maximum Likelihood Estimation technique [23] that seeks to provide a measure of fitting relative to distinct probability models by estimating parameters that maximize their Likelihood function.

The corrected Akaike Information Criterion, (correction for overfitting) is defined as $AIC_c = AIC + \frac{2k(k+1)}{n-k-1}$, where n is the sample size, k is the number of parameters, and $AIC = 2k - 2\text{LogLikelihood}(\theta)$, where θ represents the parameters to be estimated for a given model.

Let X_1, X_2, \dots, X_n be a set of continuous random variables with joint density function $f_\theta(X)$ depending on the parameters θ . The *Likelihood* function $L(\theta) = f_\theta(x_1, x_2, \dots, x_n)$, sometimes written as $L(\theta|x)$, is the joint probability distribution $f_\theta(x_1, x_2, \dots, x_n)$ with parameters θ of the set of n random variables evaluated at the observed values from the sample. The *LogLikelihood* represents the natural log of the *Likelihood* function, which is often preferred as it simplifies the calculations of critical values.

Since there is no prior knowledge of the underlying distribution of doublet return times, the AICc—by means of estimating the parameters that provide the largest plausibility for obtaining the observed values for several probability models—provides a point of comparison among the probability models that the samples are most likely to come from, serving as a means for model selection. Some of the models tested face-to-face in this sense include the Gamma, Weibull, Exponential, LogNormal, GLog, Johnson Su, Johnson Sl, Gaussian, and Normal 2 & 3 Mixture probability densities.

Akaike [23] reformulates the maximization of the *LogLikelihood* function by working with its negative value (minimization of the *LogLikelihood* function), in such case, lower values of AICc denote better model fits.

Since the AICc only provides a ranking among different types of distributions and does not warn for poorly fitted models, a Goodness-of-fit test for the model with the lowest AICc complements this part of the model selection technique, ensuring that the best-ranked model represents a good fit.

3.2 Sample size guidelines

To construct guidelines on the minimum and the maximum number of return times to consider in the distribution fit analysis, we performed simulations (with 5,000 trials at different sample sizes in the range from 5 to 5,000) by random sampling from an underlying distribution and obtained the number of times a given distribution was the best fit in the AICc sense.

The two-parameter Weibull distribution (α, β)

$$f(x; \alpha, \beta) = \frac{\beta}{\alpha} \left(\frac{x}{\alpha}\right)^{\beta-1} e^{-\left(\frac{x}{\alpha}\right)^\beta}; \quad \text{for } \alpha, \beta > 0; x \geq 0,$$

where α and β are the scale and shape parameters, respectively, was found to be the most robust at small sample sizes as it required the smallest sample size (n) to be identified as the best fit most of the trials. For instance, Fig. 9 in the Appendix section shows that at least $n \approx 6$ samples were required to achieve $\sim 30\text{-}60\%$ success rate for several parameter values of α and β , compared with Gamma ($\lambda = 4$, $scale = 1$) with at least $n \approx 26$ with $\sim 32\%$ success rate, Gaussian ($\mu = 100$, $\sigma^2 = 30^2$) with at least $n \approx 130$ with $\sim 40\%$ success rate, Normal 3 Mixture with at least $n \approx 120$ with $\sim 25\%$ success rate, and Normal 2 Mixture with at least $n \approx 38$ with $\sim 37\%$ success rate as shown in Fig. 10 in the Appendix section. Due to the high robustness of the Weibull distribution at small sample sizes that we observed with simulations, it is not surprising that the Weibull distribution is widely applied in reliability tests [24, 25, 26], which are often hampered with small sample sizes.

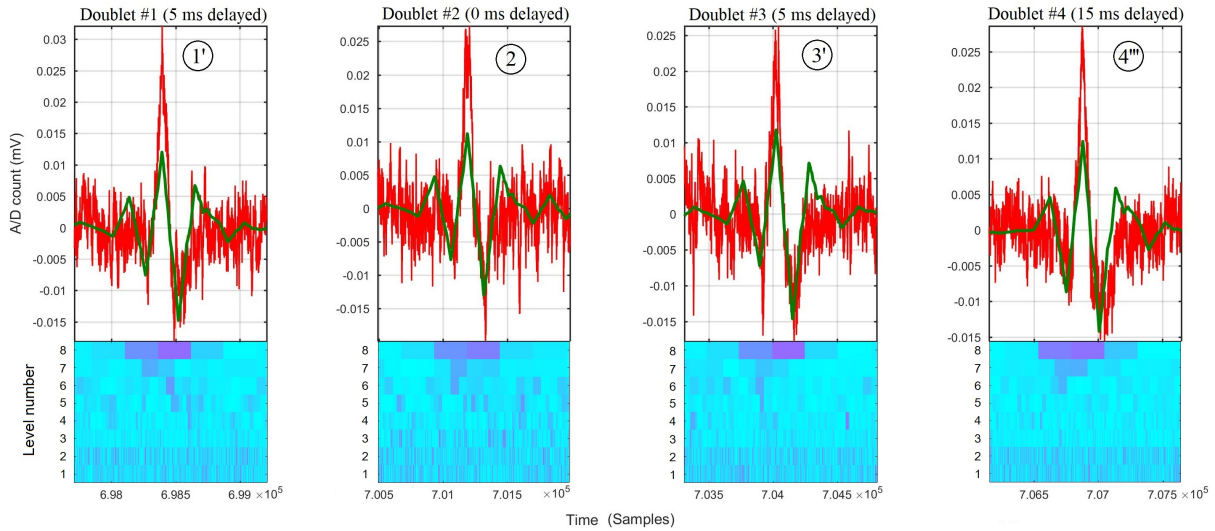


FIGURE 6

Zoom around the four bursts of raw thoracic sEMG trace (in red) of subject #4 first shown in Fig. 2, superimposed with a delayed 8-level sub-signal (in green); their respective scalogram is shown below: The D8 doublets #1', #2, #3', and #4''' correspond to time-shifting the DWT at the Pareto front time delays of 5 ms, 0 ms, 5 ms, and 15 ms, respectively, obtained for each doublet by the expert system.

As regard to the maximum number of return times to consider in the distribution fitting analysis, a stopping rule can be determined when the AICc approaches a minimum value, meaning that the percentage change of AICc approaches zero as the sample size increases as shown in Figs. 11 and 12 in the Appendix section for two different individuals.

4 Results

4.1 Results: waveform matching

Although the time locations in the first three doublets of Fig. 2 were lost in the process to retrieve the ρ -wave of D8 doublet #4, as shown in Fig. 4, the expert system recovered and found the Pareto time localization of the other doublets at each predefined delay.

With the predefined delay values inputted in the expert system for waveform matching (0, 5, 10, and 15 ms), the delays $\theta_1 = 5$ ms for doublet #1', $\theta_0 = 0$ ms for doublet #2, $\theta_1 = 5$ ms for doublet #3', and $\theta_3 = 15$ ms for doublet #4''' have been found by the expert system to be on the Pareto front (see Fig. 6).

Furthermore, the scalograms of Fig. 6, that are obtained for each burst by their respective delay times found by the expert system, also show that the $\pi\kappa\rho\sigma\tau$ wave is constructed by mainly two relatively high and successive D8 coefficients (each spanning 60 ms), correlating with our

description of a high incidence of multiple neurons firing double “exceptional” spikes [14] within each burst.

4.2 Results: theoretical probability distributions

Among the 24 paraspinal signals, the Weibull distribution was found to be the best probability fit in the AICc sense among the subjects with quadriplegia as shown in Figs. 8a and 8b; whereas the Normal 2 & 3 Mixtures were prevalently the best fit among control subjects as shown in Figs. 7a, 7b, 7c, and 7d. These results are summarized in Table 1 in the Appendix section along with the parameter estimates for the best-fitted distributions.

4.3 Results: control versus quadriplegic subjects

The contrasting difference in the results of the present study between quadriplegic and control patients, namely in their probability distributions of doublet return times and sample sizes, points to “doublets” becoming more prevalent (and with multimodal return times) in healthy neuromuscular systems than unhealthy ones. Furthermore, the more predominant rhythmic synchronization of neurophysiological activity of healthy subjects is consistent with the hypothesis that coherence at a distance is an indication of the nervous system able to coordinate the activity of many muscles [1, 9, 10].

5 Discussion

5.1 Neurophysiological personality

We hypothesize that the rhythmic bursts represent a synchronization of multiple MUs firing exceptional doublets, and that there is a probable connection between them and the dynamical system theory of the return time of rare events [27, 28] and the Generalized Extreme Value (GEV) theory of such rare events [20, 29] and the neurophysiological studies by Piotrkiewicz [14]. In the last-mentioned studies, double-firing motor units classified as single, repetitive, and exceptional doublets, constituted a small percentage (9.5%) of recorded neuronal discharges and were considered as “unusual” discharges, whereas the exceptional type was even more unusual (~1%).

In this CPG entrainment technique, those doublets deemed *exceptional* can be reproduced at will, in contrast to the studies by Piotrkiewicz [14], where the volunteers were not trained to evoke doublets.

Distribution fitting of return times among healthy subjects

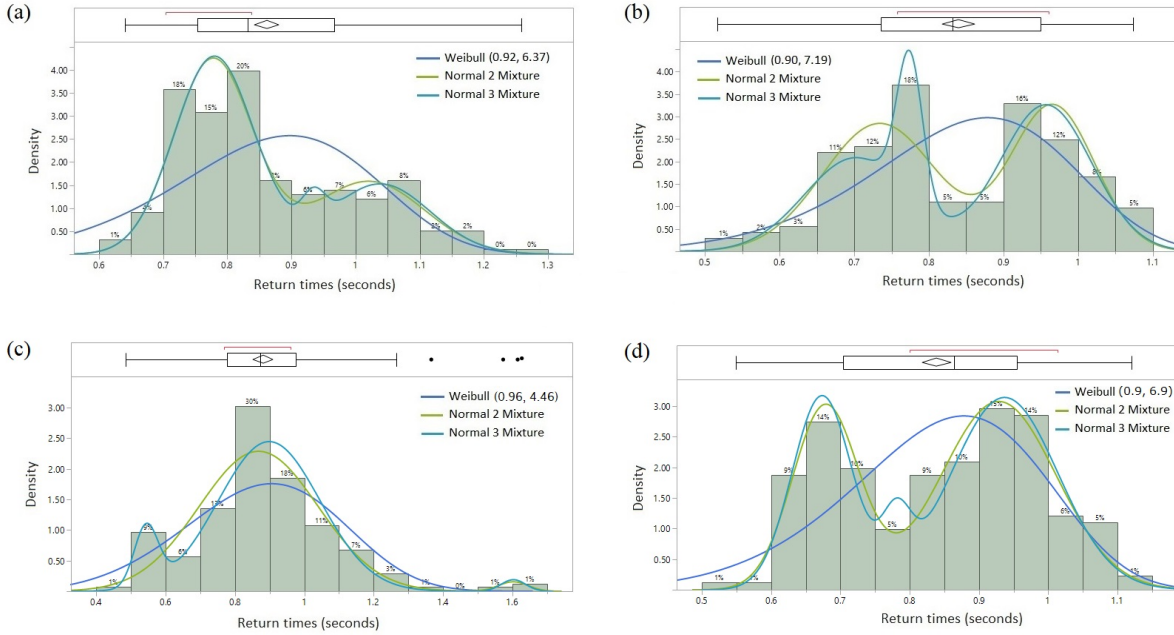


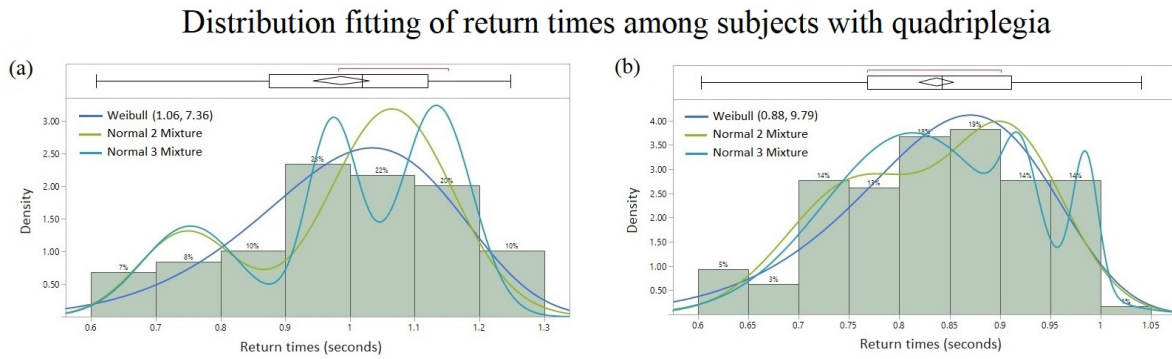
FIGURE 7

Probability distribution fitting of return times from healthy subjects: (a) lumbar spine signal of Subject #6 with Normal 2 Mixture ($AICc = -272.53$), Normal 3 Mixture ($AICc = -265.98$), and Weibull ($AICc = -199.41$), (b) sacral signal of Subject #8 with Normal 3 Mixture ($AICc = -187.61$), Normal 2 Mixture ($AICc = -186.31$), and Weibull ($AICc = -167.78$), (c) cervical signal of Subject #5 with Normal 3 Mixture ($AICc = -95.26$), Normal 2 Mixture ($AICc = -87.19$), and Weibull ($AICc = -51.56$), and (d) cervical signal of Subject #8 with Normal 2 Mixture ($AICc = -236.16$), Normal 3 Mixture ($AICc = -230.59$), and Weibull ($AICc = -194.86$). Lower $AICc$ values indicate a better distribution fit.

5.2 Heart Rate Variability and Bursting Rate Variability

Similar to the normal resting heart rate range from 60 to 100 beats per minute [30], here the observed doublet return time rate is between 60 and 88 cycles per minute, which indicates a possible connection between HRV and BRV.

It might be argued that the repetitive sEMG bursts represent a cardiac noise artifact. However, if we were observing an electrocardiographic (ECG) artifact in the sEMG traces, then we would expect the ECG artifact to have similar time parameters as those observed in a clinical ECG. Table 3.1 in [31] shows that the typical P-wave, QRS-complex, and corrected QT-interval durations for a healthy male adult have normal values and limits of 110 ± 20 ms, 100 ± 20 ms, and 400 ± 40 ms respectively. Thus, a typical cardiac PQRST-wave duration would span a total of 510 ± 60 ms. Furthermore, Fig. 3 of [30] shows a textbook example of an ECG cycle in normal conditions with a total duration of ~ 570 ms.

**FIGURE 8**

Probability distribution fitting of return times from patients with quadriplegia: (a) Thoracic signal of Subject #2 with Weibull ($AICc = -45.69$), Normal 2 Mixture ($AICc = -41.36$), and Normal 3 Mixture ($AICc = -35.46$), (b) lumbar spine signal of Subject #1 with Weibull ($AICc = -231.43$), Normal 2 Mixture ($AICc = -227.36$), and Normal 3 Mixture ($AICc = -227.12$). Lower $AICc$ values indicate a better distribution fit.

In this sEMG phenomenon, the $\pi\kappa\rho\sigma\tau$ -wave and the $\kappa\rho\sigma$ -complex span shorter durations of ~ 125 ms and ~ 62.5 ms resp., compared with the equivalent cardiac PQRST-wave and QRS-complex durations of 510 ± 60 ms and 110 ± 20 ms resp.

Besides the difference in total wavelength between the cardiac cycle and “doublets,” it is worth stressing that here variability does not appear to occur *within* the doublet (peak-dip intra-doublet time), but rather on the outside (inter-doublet return time), as the $\pi\kappa\rho\sigma\tau$ wave duration of ~ 125 ms appears to be prevalingly fixed among doublets. This is unlike HRV, where a considerable amount of variability occurs among waves within same cardiac cycles (e.g. QT prolongation [32]).

To further exemplify the difference between a pure ECG trace and the $\pi\kappa\rho\sigma\tau$ wave found here in the sEMG traces, studies show that the return time distributions of R-waves in ECG recordings have been found Erlang in normal subjects, and a weighted average of Erlang with a second distribution (e.g. Weibull) in patients with arrhythmia [33, 34].

5.3 Off-line and On-line BRV for Biofeedback Applications

For biofeedback applications, our objective is to help quadriplegic patients recover some motor control by learning how to evoke more doublet oscillations with return time distribution deviating from Weibull towards normal mixtures.

Another objective is to conduct *on-line* assessments by means of implementing the complete *off-line* technique described in section 2.5 with real-time DWT [35]. For real-time muscle per-

formance evaluations, it would reinforce the training process to help increase the number of synchronized motor units, resulting in stronger muscle contractions [36].

The proposed technique is not restricted to only paraspinal muscles as it may span the evaluation of the neuromuscular system to a greater extent—for instance, to assess rhythmic involuntary contractions such as tremors, or uterine contractions in pregnancy. In the former, it could provide feedback to therapies in the field. In the latter, it would monitor the return times of uterine EMG bursts to potentially warn for signs of imminent, false, or preterm labor [37].

6 Conclusions

The major contribution in this paper is the identification of a new neurophysiological phenomenon—the *Bursting Rate Variability* that bears some resemblance to *Heart Rate Variability*, but that still differs from it in several respects. BRV is based on recursively shifting the Daubechies 3 wavelet transform of the raw electromyographic signal to successively exhibit perfect MUAP doublets at the D8 level.

The presence of such *doublets* in the sEMG signal has been conjectured to reveal coordination of muscle masses at a distance to achieve a higher hierarchy level movement. The return time statistic of the doublets developed here adds some quantitative insights to this observation, with *Weibull to normal mixture* distribution a possible indication of a quadriplegic subject recovering some motor control.

Furthermore, from this rhythmic and highly synchronized firing pattern, the neuronal connection *structure learning* is contemplated. Whereas cardiology applications, while plausible, remain to be assessed by including electrocardiogram monitoring to our protocol while recording sEMG activity simultaneously.

Finally, from a theoretical viewpoint, this research is related to the statistic of return time of a dynamical system to some subset of its state space. The more recent Generalized Extreme Value (GEV) theory, which proceeds from the statistic of the extreme value of an observable (e.g., a sEMG signal) rather than the return time of such events, could offer an alternative way to look at the same phenomenon.

7 Appendix

Acknowledgment

Dr. Fariba Ariaei's discussions on the return time distributions in HRV dynamics are gratefully acknowledged.

TABLE 1
Distribution Fit Analysis: Enhanced-Matched Doublets

Type of Subject	Subject #	Signal	Sample Size ^a	Mean ^b			Standard Deviation ^b			Best Fit ^c	Fitted Parameter Estimates			Goodness of Fit Test ^d
				Experimental	Theoretical	%error	Experimental	Theoretical	%error		Type	Parameter	Estimate	
Quad	1	Lumbar	131	0.83679	0.83671	0.00956	0.09978	0.10269	2.83377	Weibull	Scale	α	0.8802891	0.25
	2	Thoracic	60	0.98695	0.98959	0.26678	0.16850	0.15875	6.14173	Weibull	Shape	β	9.7908324	
Control	3	Cervical	297	0.79074	0.79074	0.00073	0.07175	0.06955	3.16109	Normal 2 Mixture	Location	μ_1	0.7582432	0.4727
											Location	μ_2	0.7990039	
		Dispersion	σ_1	0.1110593										
		Dispersion	σ_2	0.0509513										
		Probability	π_1	0.2026008										
		Probability	π_2	0.7973992										
	Thoracic	142	0.80970	0.80969	0.00119	0.05908	0.05819	1.52254	Normal 3 Mixture	Location	μ_1	0.7433704	0.4605	
										Location	μ_2	0.8374764		
										Location	μ_3	1.0203324		
										Dispersion	σ_1	0.0261000		
										Dispersion	σ_2	0.0302071		
										Dispersion	σ_3	0.0200882		
	Probability	π_1	0.3226223											
	Probability	π_2	0.6632976											
	Probability	π_3	0.0140801											
	Lumbar	183	0.68051	0.68050	0.00090	0.08635	0.08508	1.49601	Normal 3 Mixture	Location	μ_1	0.4872577	0.4652	
Location										μ_2	0.6821022			
Location										μ_3	0.9507306			
Dispersion										σ_1	0.0116068			
Dispersion										σ_2	0.0720919			
Dispersion										σ_3	0.0171289			
Probability	π_1	0.0299150												
Probability	π_2	0.9543367												
Probability	π_3	0.0157483												
Sacral	257	0.71217	0.71215	0.00245	0.09731	0.09614	1.21324	Normal 2 Mixture	Location	μ_1	0.7089589	0.4707		
									Location	μ_2	1.1198451			
									Dispersion	σ_1	0.0894392			
									Dispersion	σ_2	0.0236820			
									Probability	π_1	0.9922274			
									Probability	π_2	0.0077726			
4	Cervical	1,055	0.69755	0.69756	0.00088	0.04814	0.04753	1.28368	Normal 3 Mixture	Location	μ_1	0.5211892	0.4855	
										Location	μ_2	0.7008101		
										Location	μ_3	0.6990413		
										Dispersion	σ_1	0.0106487		
										Dispersion	σ_2	0.0186435		
										Dispersion	σ_3	0.0974058		
	Probability	π_1	0.0166192											
	Probability	π_2	0.8313767											
	Probability	π_3	0.1520042											
	Thoracic	1,550	0.69693	0.69693	0.00001	0.02918	0.02857	2.12082	Normal 3 Mixture	Location	μ_1	0.6663625	0.4881	
										Location	μ_2	0.6998140		
										Location	μ_3	0.6913961		
Dispersion										σ_1	0.0868998			
Dispersion										σ_2	0.0165026			
Dispersion										σ_3	0.0331125			
Probability	π_1	0.0507722												
Probability	π_2	0.8083748												
Probability	π_3	0.1408530												

Statistical comparison of inter-burst interval times between quadriplegic and control subjects with the enhanced-matched procedure.

^a Sample size is the number of doublet return times obtained with the expert system.

^b Units of mean and standard deviation are in seconds.

^c Model selection criterion was based on the lowest AICc value.

^d Goodness-of-fit tests for Weibull and Normal Mixtures were Cramer-von Mises W and Pearson's chi-squared tests respectively.

Type of Subject	Subj. #	Signal	Sample Size ^a	Mean ^b			Standard Deviation ^b			Best Fit ^c	Fitted Parameter Estimates			Goodness of Fit Test ^d
				Experimental	Theoretical	%error	Experimental	Theoretical	%error		Type	Parameter	Estimate	
													p-value	
Control	4	Lumbar	544	0.69361	0.69361	0.00050	0.03051	0.03020	1.01183	Normal 3 Mixture	Location	μ_1	0.5177513	0.4798
											Location	μ_2	0.6904178	
											Location	μ_3	0.7028590	
											Dispersion	σ_1	0.0077241	
											Dispersion	σ_2	0.0128245	
											Dispersion	σ_3	0.0356578	
	Probability	π_1	0.0091908											
	Probability	π_2	0.6063979											
	Probability	π_3	0.3844114											
	4	Sacral	265	0.71434	0.71434	0.00057	0.06640	0.06549	1.39002	Normal 3 Mixture	Location	μ_1	0.5062527	0.4711
											Location	μ_2	0.6927545	
											Location	μ_3	0.7269021	
											Dispersion	σ_1	0.0228841	
											Dispersion	σ_2	0.0117517	
											Dispersion	σ_3	0.0636254	
	Probability	π_1	0.0223802											
	Probability	π_2	0.2231447											
	Probability	π_3	0.7544751											
5	Cervical	179	0.88047	0.88045	0.00200	0.19756	0.19505	1.28573	Normal 3 Mixture	Location	μ_1	0.5437492	0.4648	
										Location	μ_2	0.8985177		
										Location	μ_3	1.6026796		
										Dispersion	σ_1	0.0343333		
										Dispersion	σ_2	0.1466987		
										Dispersion	σ_3	0.0351786		
	Probability	π_1	0.0841804											
	Probability	π_2	0.8990634											
	Probability	π_3	0.0167563											
	5	Lumbar	199	0.96771	0.96771	0.00049	0.20595	0.19590	5.12965	Normal 2 Mixture	Location	μ_1	0.9899384	0.4667
											Location	μ_2	0.9611284	
											Dispersion	σ_1	0.3388004	
Dispersion											σ_2	0.1248396		
Probability											π_1	0.2282849		
Probability											π_2	0.7717151		
5	Sacral	361	1.00026	1.00025	0.00064	0.24514	0.24297	0.89479	Normal 2 Mixture	Location	μ_1	0.9255663	0.4753	
										Location	μ_2	1.4218577		
										Dispersion	σ_1	0.1636967		
										Dispersion	σ_2	0.1782215		
										Probability	π_1	0.8495092		
										Probability	π_2	0.1504908		
6	Thoracic	285	0.87775	0.87776	0.00069	0.17501	0.17362	0.79960	Normal 2 Mixture	Location	μ_1	0.6472983	0.4721	
										Location	μ_2	0.9608497		
										Dispersion	σ_1	0.0912896		
										Dispersion	σ_2	0.1093363		
										Probability	π_1	0.2650080		
										Probability	π_2	0.7349920		
	6	Lumbar	202	0.86156	0.86156	0.00008	0.13750	0.13609	1.03961	Normal 2 Mixture	Location	μ_1	0.7772509	0.4669
											Location	μ_2	1.0203541	
											Dispersion	σ_1	0.0616972	
											Dispersion	σ_2	0.0873302	
											Probability	π_1	0.6531933	
											Probability	π_2	0.3468067	
6	Sacral	144	0.87099	0.87099	0.00034	0.13305	0.13092	1.62844	Normal 2 Mixture	Location	μ_1	0.7801195	0.4608	
										Location	μ_2	0.9987842		
										Dispersion	σ_1	0.0563301		
										Dispersion	σ_2	0.0940075		
										Probability	π_1	0.5844435		
										Probability	π_2	0.4155565		

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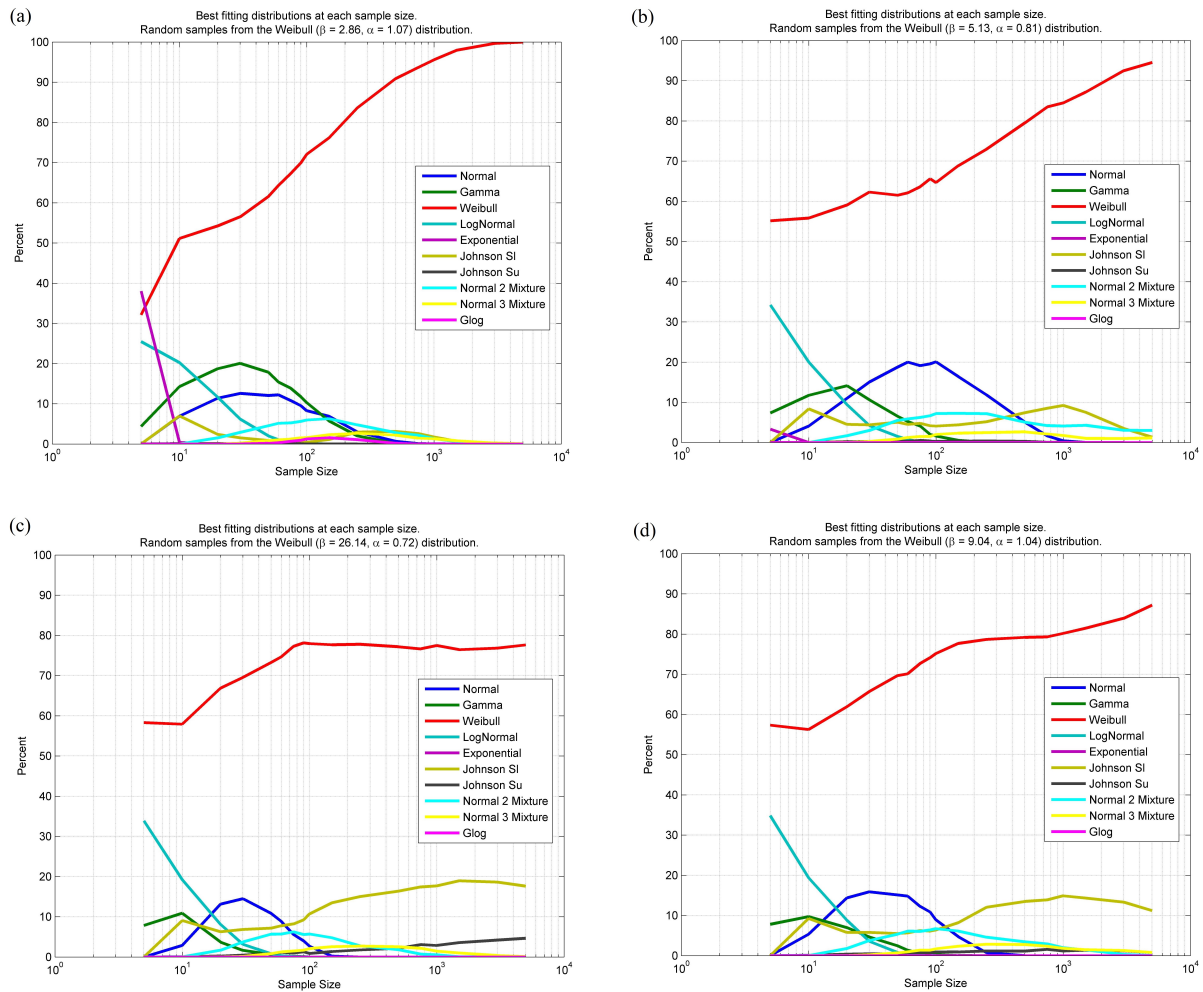


FIGURE 9

Simulation of best fitting distributions at different sample sizes obtained with the smallest AICc value by randomly sampling from (a) the Weibull ($\beta = 2.86, \alpha = 1.07$) distribution, (b) the Weibull ($\beta = 5.13, \alpha = 0.81$) distribution, (c) the Weibull ($\beta = 26.14, \alpha = 0.72$) distribution, and (d) the Weibull ($\beta = 9.04, \alpha = 1.04$) distribution.

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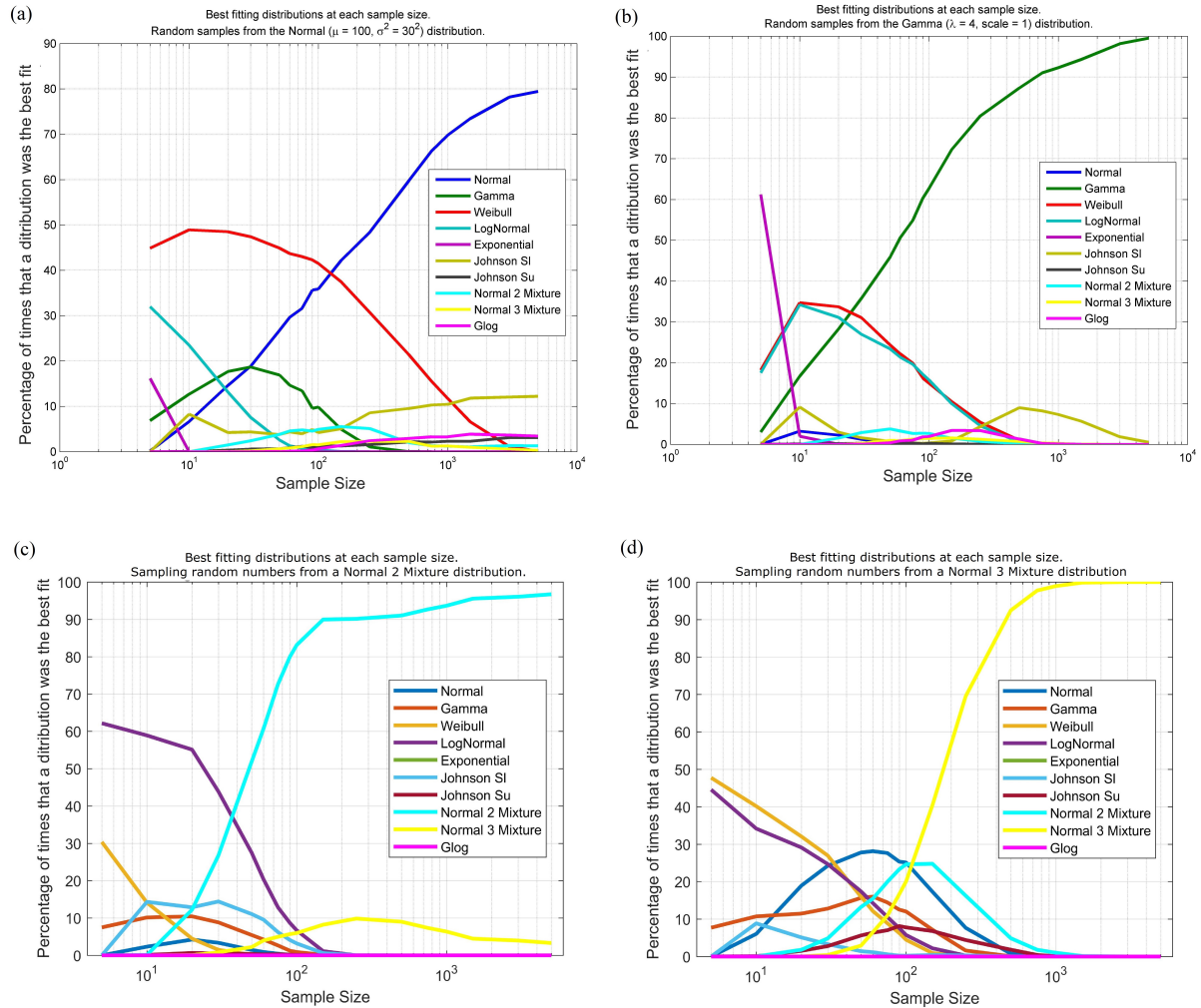


FIGURE 10

Simulation of best fitting distributions at different sample sizes obtained with the smallest AICc value by sampling random numbers from (a) the Normal ($\mu = 100, \sigma^2 = 30^2$) distribution, (b) the Gamma ($\lambda = 4, scale = 1$) distribution, (c) a Normal 2 Mixture distribution, and (d) a Normal 3 Mixture distribution.

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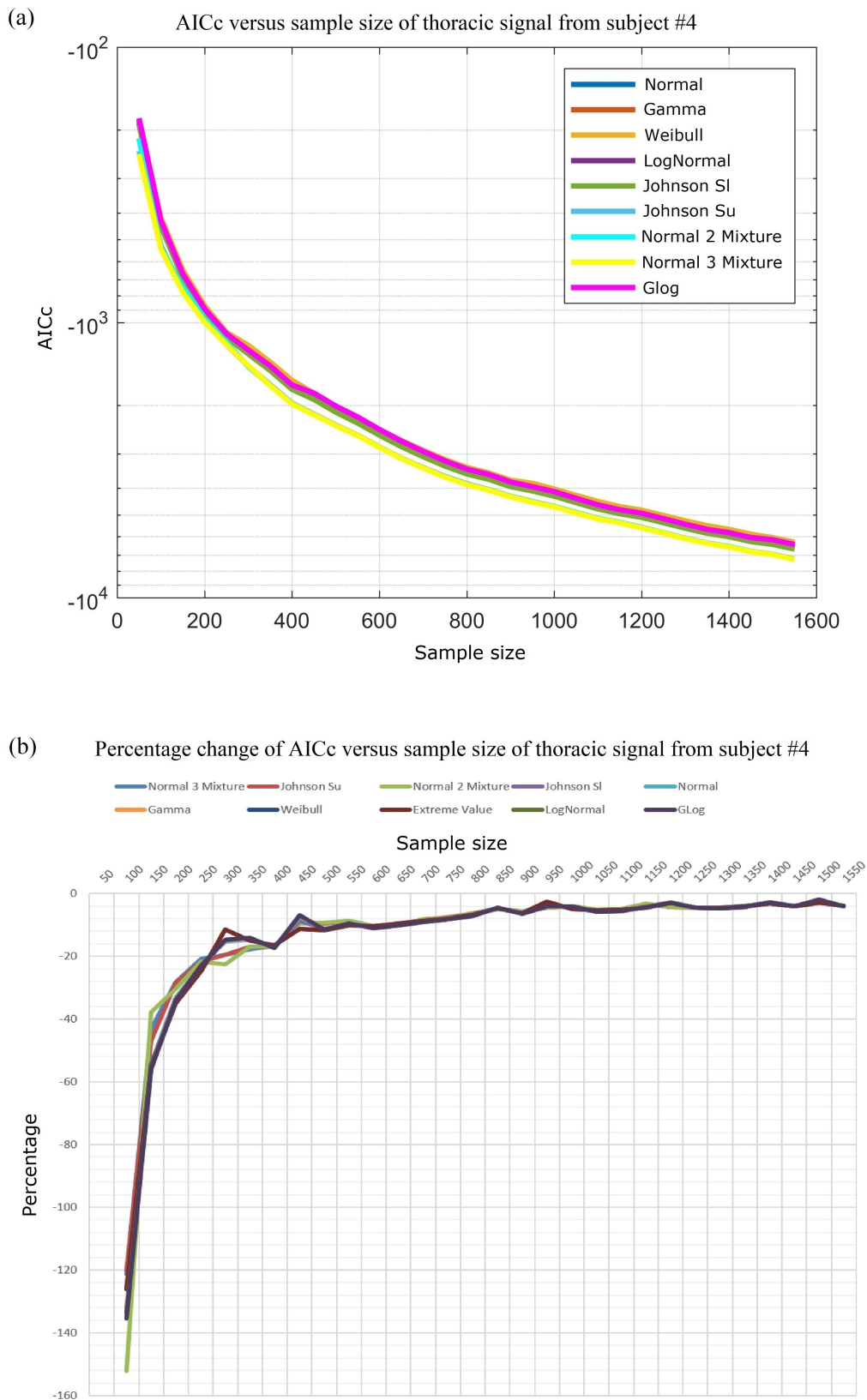


FIGURE 11

(a) AICc values versus sample size (n) to determine a stopping rule of n for thoracic signal of subject #4.

(b) Percentage change of AICc values at different sample sizes for thoracic signal of subject #4

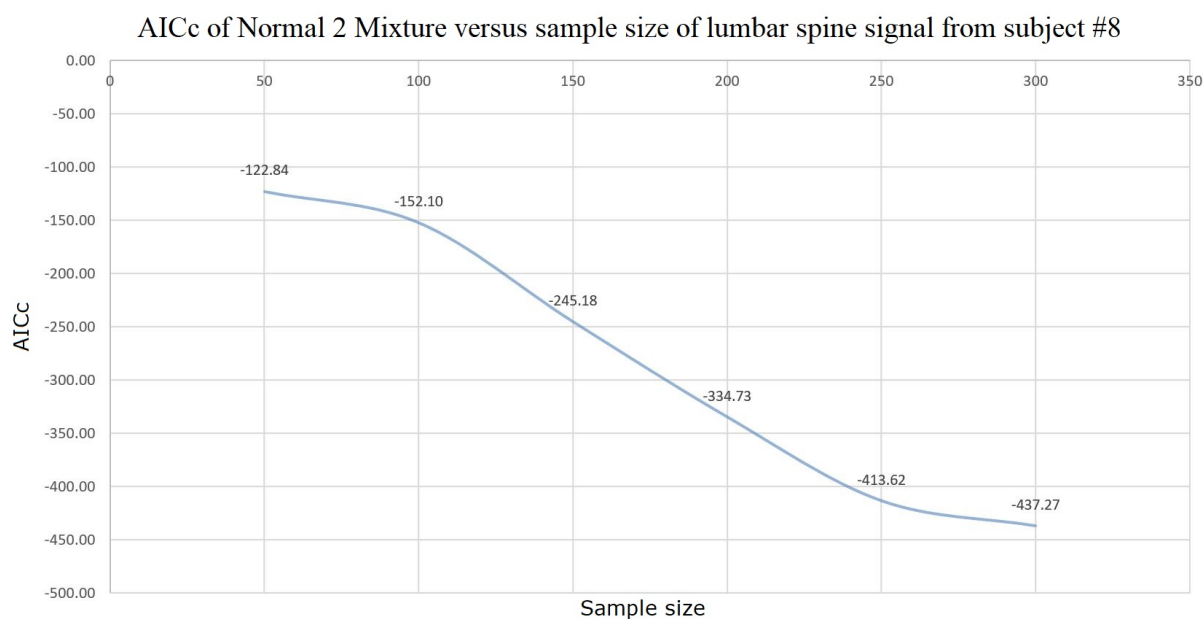


FIGURE 12

AICc values versus sample size (n) to determine a stopping rule of n for lumbar signal of subject #8.

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